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## Bayesian inference in regression models with the restricted parameter spaces

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## Tabriz, Iran (work place)







Recently, the subject that raised commonly in the framework of regression models, especially in the linear regression models and also generalized linear models, is estimating model parameters based on uncertain prior information. Taking these uncertain prior information into account when estimating the model can improve the estimation procedure in terms of accuracy, and so incorporating them has recently attached a lot of research attention. The uncertain prior information is incorporated in the model through some restrictions on parameters.

### Note

The restrictions can be linear or nonlinear functions of parameters, or even in the form of equality or inequality.

## Linear equality restrictions

### Where these restrictions come from?

- Obtained according to past data or expert opinion.
- From variable selection techniques such as AIC, CAIC and BIC.

Consider the linear equality restrictions on the model parameters that can be written as:

$$R\beta = r$$

where  $R_{q \times p}$  and  $r_{q \times 1}$  are a pre-specified matrix and vector.

## Linear equality restrictions

### What are the estimators?

- Unrestricted estimator
- Rerestricted estimator
- Shrinkage methods
  - ① Linear shrinkage estimator
  - ② Pretest estimator
  - ③ James-Stein type estimator

## Linear inequality restrictions

### Where these kind of restrictions come from?

In the real world, physical quantities show specific behaviors:

- Chemical concentrations lies between 0 and 1.
- Due to the physical nature, Some models usually have restrictions on the sign of the input variables
- In econometrics, we encounter situations where some parameters of the model should be considered non-negative or negative.
- In hyperspectral imaging, due to physical considerations, coefficient parameters should be considered non-negative.

## What to do?

From a practical point of view, it would be highly desirable to build a surrogate model which respects the same restrictions to ensure structural compatibility based on physical phenomena or the validity of scientific theories.

The linear inequality restrictions on the model:

$$L_{q \times 1} \leq H_{q \times p} \beta_{p \times 1} \leq G_{q \times 1}$$

It contains:

- linear equality restrictions
- Order restrictions
- Sign restrictions

## Bayesian inference in Linear regression model with linear inequality restrictions

The linear regression model is given by:

$$Y = \mathbf{X}\boldsymbol{\beta} + \varepsilon, \quad \varepsilon \sim N(0, \sigma^2)$$

where  $\boldsymbol{\beta} = (\beta_1, \beta_2, \dots, \beta_p)^T$  is a vector of the regression coefficients.  
Consider  $q$  linear inequality restrictions on the model parameters as:

$$K \leq H\boldsymbol{\beta} \leq G$$

where  $H$  is a matrix of size  $q \times p$  with ( $q \leq p$ ) and  $G$  and  $K$  are two vectors of length  $q$ .

## Classical inference

- Judge and Takayama (1966) used Dantzig-Cottle algorithm (IROLS)
- Escobar and Skarpness (1986, 1987) obtained the bias and efficiency of IROLS
- Ohtani (1987) obtained the MSE of IROLS

## Bayesian inference

- Geweke (1986):
  - $\pi(\beta, \sigma^2) \propto \frac{1}{\sigma} I_{(K \leq H\beta \leq G)}$
  - Method: importance sampler
  - When  $p$  increases or the posterior probability of restriction is small, the algorithm will be slow.
- Geweke (1996)
  - Method: Gibbs sampler
  - Algorithm is practical when  $H$  is invertable.
- Rodriguez-Yam et al. (2004):
  - Method: Gibbs sampler
  - Algorithm is practical when the coefficients only are upper bounded.

# Prior specification

We Partition  $H = (H_S, H_{S'})$

How?

Suppose that  $P = \{1, \dots, p\} = S \cup S'$  is a set of the indices of the matrix  $H$ .

$S = \{j : h_{ij} \in H_S; i = 1, \dots, q\} \subset P$  is a subset of the indices of the columns of  $H$  for which the block  $H_S$  is a full rank matrix.

The result

The matrix  $H_S$  is then a  $q \times q$  full rank block of  $H$ .

$H_{S'}$  is a  $q \times (p - q)$  sub-matrix including the rest of the elements of  $H$ .

## Consequence

$$\boldsymbol{\beta} = (\boldsymbol{\beta}_S^T, \boldsymbol{\beta}_{S'}^T)^T$$

Therefore:

$$K \leq (H_{S(q \times q)}, H_{S'((p-q) \times (p-q))}) \begin{pmatrix} \boldsymbol{\beta}_{S(q \times 1)} \\ \boldsymbol{\beta}_{S'((p-q) \times 1)} \end{pmatrix} \leq G$$

$$K \leq H_S \boldsymbol{\beta}_S + H_{S'} \boldsymbol{\beta}_{S'} \leq G$$

$$K - H_{S'} \boldsymbol{\beta}_{S'} \leq H_S \boldsymbol{\beta}_S \leq G - H_{S'} \boldsymbol{\beta}_{S'}$$

## Prior distributions

①  $\sigma^2 \sim IG\left(\frac{a}{2}, \frac{b}{2}\right)$

②  $\boldsymbol{\beta}_{S'} | \sigma^2 \sim N(\mu_{\boldsymbol{\beta}_{S'}}, \sigma^2 \Sigma_{\boldsymbol{\beta}_{S'}})$

③  $\boldsymbol{\beta}_S | \boldsymbol{\beta}_{S'}, \sigma^2 \sim TN(\mu_{\boldsymbol{\beta}_S}, \sigma^2 \Sigma_{\boldsymbol{\beta}_S}, T)$  where

$$T = \{\boldsymbol{\beta}_S \in \mathbb{R}^q | K - H_{S'} \boldsymbol{\beta}_{S'} \leq H_S \boldsymbol{\beta}_S \leq G - H_{S'} \boldsymbol{\beta}_{S'}\}.$$

## Posterior distributions

$$\begin{aligned}Y &= \mathbf{X}_{S;S'} \boldsymbol{\beta}_{S;S'} + \boldsymbol{\epsilon} \\&= (\mathbf{X}_S, \mathbf{X}_{S'}) \begin{pmatrix} \boldsymbol{\beta}_S \\ \boldsymbol{\beta}_{S'} \end{pmatrix} + \boldsymbol{\epsilon} \\&= \mathbf{X}_S \boldsymbol{\beta}_S + \mathbf{X}_{S'} \boldsymbol{\beta}_{S'} + \boldsymbol{\epsilon}\end{aligned}$$

The posterior distributions will be:

- $\sigma^2 | Y, \mathbf{X}_{S;S'} \sim IG(\tilde{v}, \tilde{\eta})$   
where

$$\tilde{v} = \frac{n+a}{2}$$

$$\tilde{\eta} = \frac{1}{2} (b + Y^T Y + \mu_{\beta_{S'}}^T \Sigma_{\beta_{S'}}^{-1} \mu_{\beta_{S'}} + \mu_{\beta_S}^T \Sigma_{\beta_S}^{-1} \mu_{\beta_S} - W^T \tilde{\Sigma}_{\beta_S} W - \tilde{\mu}_{\beta_{S'}} \tilde{\Sigma}_{\beta_{S'}}^{-1} \tilde{\mu}_{\beta_{S'}})$$

## Posterior distributions

- $\boldsymbol{\beta}_{S'} | \sigma^2, Y, \mathbf{X}_{S;S'} \sim N(\tilde{\mu}_{\beta_{S'}}, \sigma^2 \tilde{\Sigma}_{\beta_{S'}})$  where

$$\tilde{\Sigma}_{\beta_{S'}} = \mathbf{X}_{S'}^T \mathbf{X}_{S'} + \Sigma_{\beta_{S'}}^{-1} - (\mathbf{X}_S^T \mathbf{X}_{S'})^T \tilde{\Sigma}_{\beta_S} (\mathbf{X}_S^T \mathbf{X}_{S'})$$

$$\tilde{\mu}_{\beta_{S'}} = \tilde{\Sigma}_{\beta_{S'}} (\mathbf{X}_{S'}^T Y + \Sigma_{\beta_{S'}}^{-1} \mu_{\beta_{S'}} - (\mathbf{X}_S^T \mathbf{X}_{S'})^T \tilde{\Sigma}_{\beta_S} W)$$

- $\boldsymbol{\beta}_S | \boldsymbol{\beta}_{S'}, \sigma^2, Y, \mathbf{X}_{S;S'} \sim TN(\tilde{\mu}_{\beta_S}, \sigma^2 \tilde{\Sigma}_{\beta_S}, T)$  where

$$\tilde{\Sigma}_{\beta_S} = \mathbf{X}_S^T \mathbf{X}_S + \Sigma_{\beta_S}^{-1}$$

$$\tilde{\mu}_{\beta_S} = \tilde{\Sigma}_{\beta_S} (W - \mathbf{X}_S^T \mathbf{X}_{S'} \boldsymbol{\beta}_{S'})$$

$$W = \mathbf{X}_S^T Y + \Sigma_{\beta_S}^{-1} \mu_{\beta_S}$$

## Proposed algorithm (Bevrani Kamary Seifollahi (BKS) algorithm)

- Initialize  $\beta^{(0)}$  and  $\sigma^{2(0)}$ ,
- Update  $\beta^{(t)}$  and  $\sigma^{2(t)}$  for  $t : 1, 2, \dots$  by the below steps:
  - Step 1: Draw  $\sigma^{2(t)} | Y, \mathbf{X}_{S;S'} \sim IG(\tilde{v}, \tilde{\eta})$ ,
  - Step 2: Draw  $\beta_{S'}^{(t)} | \sigma^{2(t)}, Y, \mathbf{X}_{S;S'} \sim N(\tilde{\mu}_{\beta_{S'}}, \sigma^{2(t)} \tilde{\Sigma}_{\beta_{S'}})$ ,
  - Step 3: Calculate the restriction inequality bounds and the posterior mean of  $\beta_S$  from the following equation:

$$T^{(t)} = \{\beta_S \in \mathbb{R}^q | K - H_{S'} \beta_{S'}^{(t)} \leq H_S \beta_S \leq G - H_{S'} \beta_{S'}^{(t)}\},$$

$$\tilde{\mu}_{\beta_S} = \tilde{\Sigma}_{\beta_S} (W - \mathbf{X}_S^T \mathbf{X}_{S'} \beta_{S'}^{(t)})$$

- Step 4: Draw  $\beta_S^{(t)} | \beta_{S'}^{(t)}, \sigma^{2(t)}, Y, \mathbf{X}_{S;S'} \sim TN(\tilde{\mu}_{\beta_S}, \sigma^{2(t)} \tilde{\Sigma}_{\beta_S}, T^{(t)})$ .

Consider  $y_i = \beta_1 + \beta_2 x_{i1} + \beta_3 x_{i2} + \beta_4 x_{i3} + \beta_5 x_{i4} + \varepsilon_i \quad i : 1, 2, \dots, 20$

$$\varepsilon_i \sim N(0, 1)$$

$$x_{ij} \sim N(0, 1)$$

$$\boldsymbol{\beta} = (-0.5, 1, -2, 3, 4)^T$$

Restriction 1	Restriction 2
$\beta_2 + \beta_3 \leq -0.5$	$\beta_2 + \beta_3 \leq -0.5$
$\beta_2 + \beta_4 - \beta_5 \leq 0.2$	$\beta_3 \leq -1.5$
$\beta_3 + \beta_5 \leq 2.2$	$\beta_4 \geq 2$
BKS	BKS and Geweke

$$H^{[1]} = \begin{pmatrix} 0 & 1 & 1 & 0 & 0 \\ 0 & 1 & 0 & 1 & -1 \\ 0 & 0 & 1 & 0 & 1 \end{pmatrix}, \quad G^{[1]} = \begin{pmatrix} -0.5 \\ 0.2 \\ 2.2 \end{pmatrix}$$

$$H^{[2]} = \begin{pmatrix} 0 & 1 & 1 & 0 & 0 \\ 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 0 & -1 & 0 \end{pmatrix}, \quad G^{[2]} = \begin{pmatrix} -0.5 \\ -1.5 \\ -2 \end{pmatrix}$$

$$H_G = \begin{pmatrix} 1 & 0 & 0 & 0 & 0 \\ 0 & 1 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & -1 & 0 \\ 0 & 0 & 0 & 0 & 1 \end{pmatrix}, \quad G_G = \begin{pmatrix} +\infty \\ -0.5 \\ -1.5 \\ -2 \\ +\infty \end{pmatrix}$$

$$H_S^{[1]} = \begin{pmatrix} 1 & 0 & 0 \\ 0 & 1 & -1 \\ 1 & 0 & 1 \end{pmatrix}, H_{S'}^{[1]} = \begin{pmatrix} 0 & 1 \\ 0 & 1 \\ 0 & 0 \end{pmatrix}, H_S^{[2]} = \begin{pmatrix} 1 & 1 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & -1 \end{pmatrix}, H_{S'}^{[2]} = \mathbf{0}_{3 \times 2}$$

$$\boldsymbol{\beta}_S^{[1]} = (\beta_3, \beta_4, \beta_5)^T, \boldsymbol{\beta}_{S'}^{[1]} = (\beta_1, \beta_2)^T, \boldsymbol{\beta}_S^{[2]} = (\beta_2, \beta_3, \beta_4)^T, \boldsymbol{\beta}_{S'}^{[2]} = (\beta_1, \beta_5)^T$$

$$\mathbf{X}_S^{[1]} = (X_2, X_3, X_4), \mathbf{X}_{S'}^{[1]} = (\mathbf{1}_n, X_1), \mathbf{X}_S^{[2]} = (X_1, X_2, X_3), \mathbf{X}_{S'}^{[2]} = (\mathbf{1}_n, X_4)$$

## Prior distributions:

$$\sigma^2 \sim IG(3, 1)$$

$$\boldsymbol{\beta}_{S'} | \sigma^2 \sim N_2(\mu_{S'}, \sigma^2 (\mathbf{X}_{S'}^T \mathbf{X}_{S'})^{-1})$$

$$\boldsymbol{\beta}_S | \boldsymbol{\beta}_{S'}, \sigma^2 \sim TN_3(\mu_S, \sigma^2 (\mathbf{X}_S^T \mathbf{X}_S)^{-1}, M)$$

where

$$\hat{\boldsymbol{\beta}}_{OLS} = \mu = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{Y}$$

$$M = \{\boldsymbol{\beta}_S \in \mathbb{R}^3 | H_S \boldsymbol{\beta}_S \leq G - H_{S'} \boldsymbol{\beta}_{S'}\}.$$

Prior distribution for Geweke method:

$$\boldsymbol{\beta} | \sigma^2 \sim N_5(\mu, \sigma^2 (\mathbf{X}^T \mathbf{X})^{-1}). I_{(H_G \boldsymbol{\beta} \leq G_G)}$$

First, we generate  $10^4$  data from the posterior distribution by the proposed algorithm.

	Parameters	stationary	p-value	Halfwidth	Mean	Halfwidth
		test		test		
Rest 1.	$\beta_1$	passed	0.356	passed	-0.329	0.002
	$\beta_2$	passed	0.568	passed	1.195	0.002
	$\beta_3$	passed	0.543	passed	-1.996	0.007
	$\beta_4$	passed	0.949	passed	2.794	0.002
	$\beta_5$	passed	0.260	passed	3.931	0.003
Rest 2.	$\beta_1$	passed	0.293	passed	-0.329	0.002
	$\beta_2$	passed	0.210	passed	1.192	0.001
	$\beta_3$	passed	0.822	passed	-1.994	0.005
	$\beta_4$	passed	0.126	passed	2.910	0.002
	$\beta_5$	passed	0.550	passed	3.820	0.002

Table : The Heidelberg and Welch diagnostic results of coefficients based on generated data.

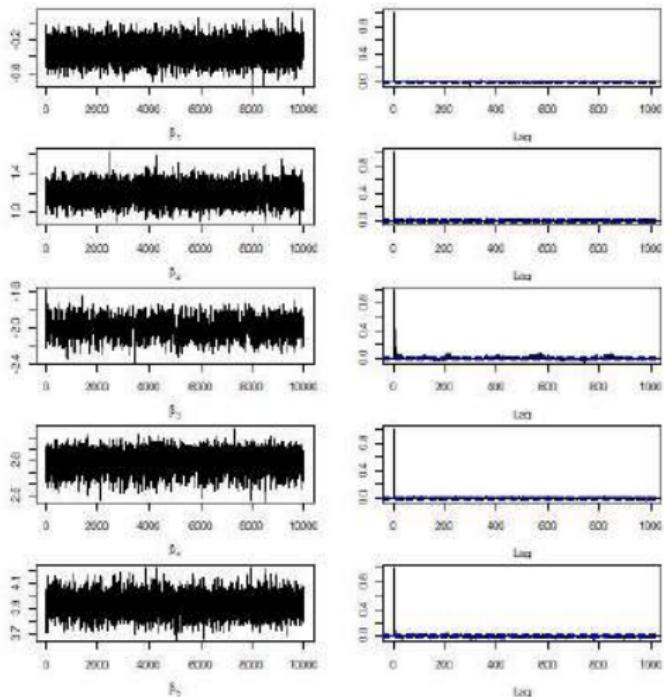


Figure : Trace and the sample ACF plots of generated data in Rest 1.

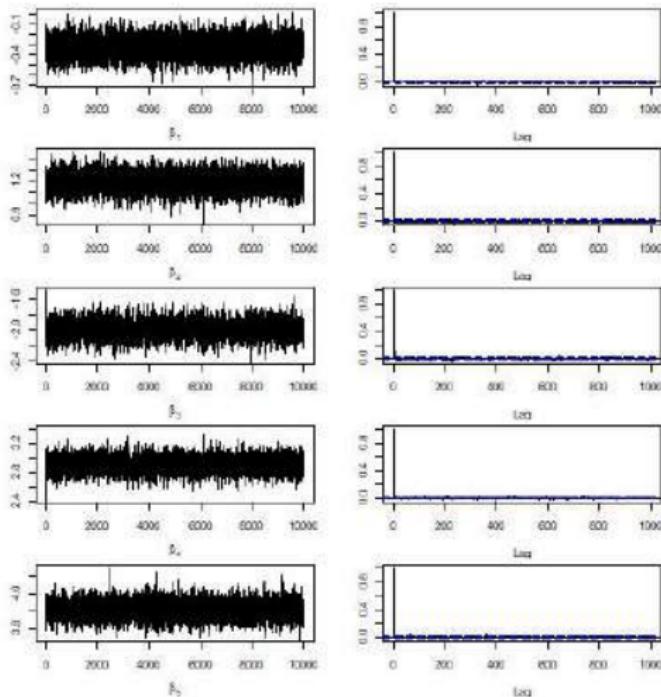


Figure : Trace and the sample ACF plots of generated data in Rest 2.

Gibbs path of length  $10^4$  with 2000 data as burn-in and 2000 replication.  
The loss function is quadratic function. The criteria for comparing estimators:

$$\text{MSE}(\hat{\boldsymbol{\beta}}) = \frac{1}{2000} \sum_{k=1}^{2000} \sum_{j=1}^p (\hat{\beta}_{jk} - \beta_j^{real})^2$$

$$\text{RE} = \frac{\text{MSE}(\hat{\boldsymbol{\beta}}_{UN})}{\text{MSE}(\hat{\boldsymbol{\beta}}_{RE})}$$

		Restriction 1.	Restriction 2.	
	Unrestricted	BKS-method	BKS-method	Geweke's method
$\beta_1$	$\hat{\beta}_1$	-0.5019	-0.5016	-0.4984
	SE	0.2610	0.2545	0.2527
$\beta_2$	$\hat{\beta}_2$	0.9960	1.0013	0.9853
	SE	0.2678	0.2614	0.2387
$\beta_3$	$\hat{\beta}_3$	-1.9990	-2.0002	-2.0301
	SE	0.2718	0.2019	0.2205
$\beta_4$	$\hat{\beta}_4$	3.0031	2.9983	3.0018
	SE	0.2655	0.2216	0.2630
$\beta_5$	$\hat{\beta}_5$	3.9957	4.0025	4.0009
	SE	0.2793	0.1831	0.2580
<i>MSE</i>		0.0724	0.0557	0.0602
<i>RE</i>		-	1.30	1.20
time		-	-	66.54
				82.64

Table : Bayesian estimations and SE of parameters with different methods

We considered the illustration provided by Pindyck and Rubinfeld (1981) on page 44 about undergraduates at the University of Michigan.

$$y_i = \beta_1 + \beta_2 s_i r_i + \beta_3 (1 - s_i) r_i + \beta_4 s_i d_i + \beta_5 (1 - s_i) d_i + \varepsilon_i \quad i : 1, 2, \dots, 32$$

where  $y_i$  denotes rent paid per person,  $d_i$  distance from campus in blocks,  $r_i$  number of rooms per person,  $s_i$  is a dummy variable representing gender (one for male and zero for female).

The restrictions considered on the  $\beta$ 's are

$$\beta_2 \geq 0, \beta_3 \geq 0, \beta_4 \leq 0, \beta_5 \leq 0$$

$$H_{BKS} = \begin{pmatrix} 0 & -1 & 0 & 0 & 0 \\ 0 & 0 & -1 & 0 & 0 \\ 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 1 \end{pmatrix}, \quad G_{BKS} = \begin{pmatrix} 0 \\ 0 \\ 0 \\ 0 \end{pmatrix}.$$

$$H_G = \begin{pmatrix} 1 & 0 & 0 & 0 & 0 \\ 0 & -1 & 0 & 0 & 0 \\ 0 & 0 & -1 & 0 & 0 \\ 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 1 \end{pmatrix}, \quad G_G = \begin{pmatrix} +\infty \\ 0 \\ 0 \\ 0 \\ 0 \end{pmatrix}$$

$$H_S = \begin{pmatrix} -1 & 0 & 0 & 0 \\ 0 & -1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{pmatrix}, \quad H_{S'} = \begin{pmatrix} 0 \\ 0 \\ 0 \\ 0 \end{pmatrix}.$$

Hyperparameters:

$$a = b = 0.001, \quad \mu_S = \begin{pmatrix} 130.0 \\ 123.0 \\ 0.0 \\ -1.153 \end{pmatrix}, \quad \mu_{S'} = 37.63,$$
$$\Sigma_S = (X_S^T X_S)^{-1}, \quad \Sigma_{S'} = (X_{S'}^T X_{S'})^{-1}.$$

For Geweke method:

$$\mu_\beta = (37.63, 130.0, 123.0, 0.0, -1.153)^T, \quad \Sigma_\beta = (X^T X)^{-1}$$

Parameters	BKS-method	Geweke's method
$\beta_1$	37.7037 (5.1998)	37.8017 (21.9386)
$\beta_2$	134.8952 (9.8633)	134.9228 (24.4277)
$\beta_3$	122.7444 (9.7056)	122.7571 (25.3568)
$\beta_4$	-0.6447 (0.5692)	-0.6573 (0.5787)
$\beta_5$	-1.1448 (0.3872)	-1.1522 (0.3938)
$\sigma^2$	1316.165 (349.1228)	1323.586 (353.8908)

Table : Estimates and SD by using the BKS-method and Geweke's method based on  $10^4$  Gibbs paths.

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Thank You for Your Attention

