## Space-Time Monitoring of Count Data for Public Health Surveillance

Arda Vanli<sup>1</sup>, Nour Alawad<sup>1</sup>, Rupert Giroux<sup>2</sup> <sup>1</sup> Department of Industrial and Manufacturing Engineering, Florida A&M University – Florida State University College of Engineering

> <sup>2</sup> Florida Department of Transportation, State Safety Office

ENBIS 2021 Online Spring Meeting: Data Science in Process Industries 17 – 18 May 2021



FAMU-FSU College of Engineering

### Outline

- Introduction: Geographical anomaly detection and public health surveillance
- Proposed Methodology: Space-time CUSUM for trends in Poisson counts
- Simulation results
- Case study results



## **Geographical anomaly detection**

- Fast statistical anomaly detection on streaming large scale data
- Anomaly: any pattern that is different from behavior that was "expected" or "normal" based on past information
  - Anomalies are time segments (for Temporal), regions (for Geographic) or vertices/edges (for Network) that have larger incidence rates than would be observed under normal conditions
- Hypothesis testing
  - $H_0$ : Data comes from a reference pdf that is homogeneous through space
  - $H_1$ : Inhomogeneities (clusters) exist in the data pdf
- Continuous monitoring of a system
  - Sequential measurements from system, test hypothesis, detect any anomaly as quickly as possible
  - False alarm prob, misdetection prob, reference pdf, divergence from normalcy



## Public health surveillance

- Detect geographical clusters, or regions of anomalous activity: identify the nearby areal units with incident rates higher than expected (baseline) values
  - Baseline disease rates are estimated from historical data
  - Spatio-temporal surveillance: Sequentially take measurements of disease incidences from the map of areal units; test the hypothesis of no spatial clusters
  - Is pattern observed in today's map different from a pattern that was "expected"? Is the observed deviation from expected pattern result of some noteworthy event (e.g., Disease outbreak, contamination in waterways, crime counts in neighborhood, traffic crashes on roadways)?
- Timely and accurate detection of emerging geographical disease clusters is critical to devise effective epidemic containment and mitigation policies



## Public health surveillance: Covid-19 Outbreaks

## in Florida

- Weekly COVID-19 case counts observed from Feb to May 2020 in Miami, FL
  - Zip Codes are areal units
  - Expected value is found by assuming all counts are spatially and temporally homogeneous, from the first 4 weeks of data
- Identify emerging geographical clusters with higher than "expected" incidence rates
  - Determine *when* and *where* anomalously high disease rates are beginning to occur

#### Miami, FL, zip codes



#### Covid outbreaks in Miami, FL, 2020





## **Proposed method: Space-time CUSUM**

- Space-time CUSUM to detect trend-type shifts in regional Poisson count data<sup>\*</sup>
  - Enumerate a set of overlapping cylinders with circular bases over a geographic area (with varying centers and radii), and a sliding interval of time (varying heights)
    - Cylinder Z with circular base centered at c, radius r and height that correspond to the time period  $[\tau, k]$  between outbreak onset time  $\tau$  and current time k
  - Calculate the local CUSUMs for all possible cylinders, find the maximum of all local CUSUMs → most unlikely cylinder under the null hypothesis that the rate of incidents is homogeneous over the entire space
  - Estimate of the onset of outbreak (change point)
  - Space-time CUSUM assumes a hypothesized outbreak infection rate (Sonesson, 2007). Space-time Scan uses a generalized likelihood ratio test (Kuldorff, 2001)

\* Vanli, Alawad, (2020), Space-Time Surveillance of Count Data Subject to Linear Trends, *Quality and Reliability Engineering* International



Study area: Miami, FL metro area Areal units: Zip codes Search grid: Zip code centroids





## Circular spatial boundary to aggregate counts over space <u>A: set of zip</u>





#### Enumerate over *c*





#### Enumerate over *c*





#### Enumerate over *c*





#### Enumerate over *r*





#### Enumerate over *r*









### **Proposed method**

- Count data  $y_{ik}$  at subregion i and time t follows Poisson with incidence rate  $\mu_0$ . Test the spatial hypotheses:
  - $H_0$ : infection rate is  $\mu_0$  for all sub-regions

College of Engineering

- $H_1$ : infection rate of some contiguous sub-regions has shifted according to  $\mu_1^*(t) = \mu_0 + \theta^* \sqrt{\mu_0}(t \tau + 1)$ , with drift rate  $\theta^*$  at the change-point  $\tau$
- The log likelihood ratio (LLR) of data observed up to current time k, within the set A of subregions (cylinder Z) and time interval  $[\tau, k]$

$$L(c, r, \tau, k) = \sum_{t=\tau}^{k} \log \prod_{i \in A} \frac{f(y_{it} | \mu_1^*)}{f(y_{it} | \mu_0)}$$

• Maximum of LLR over all  $\tau$  to determine change-point is a local cumulative sum (CUSUM)

$$\max_{1 \le \tau \le k} L(c, r, \tau, k) \equiv T_{rc}(k) = \max\left\{0, T_{rc}(k-1) + \log \prod_{i \in A} \frac{f(y_{it}|\mu_1^*)}{f(y_{it}|\mu_0)}\right\}$$

## **Proposed method**

• Maximum of the LLR over all  $\tau$ , c and r to determine change-point, location and geographical size

 $\max_{r \in R} \max_{c \in C} \max_{1 \le \tau \le k} L(c, r, \tau, k) = \max_{r \in R} \max_{c} T_{rc}(k)$ 

– maximum of local CUSUMs,  $T_{rc}$ , defined for all r and c

- Center and size of the most likely cluster emerging at t $\{r^*, c^*\} = \underset{r \in R. c \in C}{\operatorname{argmax}} T_{rc}(t)$
- Change-point estimate

$$\hat{\tau}_{r^*,c^*} = \max_{1 \le t \le k} \{ t | T_{r^*c^*}(t) = 0 \}$$
 where

 Compare to Sonesson (2007) which considered detecting step-type sustained shifts in Poisson counts

## Simulation study

- Infection rate  $\mu_0$  is homogeneous spatially and it starts shifting at time  $t_0$  with slope  $\theta$  according to  $\mu_1 = \mu_0 + \theta \sigma_0 (t - t_0 + 1)$ 
  - Baseline rate  $\mu_0 = 1.4$
  - slope  $\theta = 0.1$
- Outbreak scales: Case A (localized) and Case B (regional)
- The set of possible radius and center values used in monitoring:
  r ∈ {0,1,2} for c ∈ {1,2,...,36}
- Results from 20 sample simulations where trend shift is introduced at  $t_0 = 15$ .













## Simulation Study

- Proposed trend-type CUSUM (T-SCUSUM) was tuned to detect slopes  $\theta^* = 0.10, 0.50$  and 1.00
- Sonesson (2007) step shift-type CUSUM (S-SCUSUM) was tuned to detect shifts of sizes  $\delta^* =$ 0.81, 2.14 and 4.00 (standard deviations)
- Methods are used to monitor outbreaks that started at time t = 30 with drifts θ = 0.10, 0.50 and 1.00
- Simulations repeated 10,000 times







Case B



# Trend and step shift type CUSUMs are designed to have approximately same ARL

	Slope	T-SCUSUM ( $\theta^*$ )			S-SCUSUM ( $\delta^*$ )		
Case	θ	0.1	0.5	1	0.81	2.14	4.00
А	0.1	4.31	6.61	7.87	4.25	6.28	6.72
	0.5	2.90	3.71	4.13	2.92	3.68	3.83
	1	2.38	2.82	2.86	2.29	2.72	2.79
В	0.1	3.17	4.63	5.63	3.20	4.77	5.36
	0.5	1.98	2.28	2.59	1.94	2.36	2.65
	1	1.59	1.65	1.80	1.47	1.67	1.88



## Change point estimates from T-SCUSUM and S-CUSUM

• Change point estimates  $\hat{\tau}$  and the improvements obtained by the use of a trend type detector over the use of a step type detector



### Case study 1: New Mexico thyroid cancer data

- Male thyroid cancer incidences in 32 counties of New Mexico between 1975 and 2016
- Baseline incidence rate estimated from 1975 to 1988
  - Rates of counties  $\mu_{0it}$  are non-homogeneous.
  - The non-homogeneities are assumed to be due to nonhomogeneous population sizes of the subregions  $\mu_{0it} = n_{it}\lambda_0$
  - $n_{it}$ : population (in 100K) in county *i* and year *t*
  - Baseline rate for entire state:  $\lambda_0$  (per 100,000 persons)

$$\lambda_0 = \frac{1}{14} \sum_{1975}^{1988} \lambda_t$$
 and  $\lambda_t = \frac{1}{32} \sum_{i=1}^{32} \frac{y_{it}}{n_{it}}$ 

• Largest scan radius includes half of the state population







## Case study 1:

- Using the T-SCUSUM (tuned for two different rates) clusters centered at Los Alamos was detected in 1994 or 1993
- Using the S-SCUSUM (tuned for two different step sizes) a cluster centered at Socorro was detected in 1995 and a cluster centered at Bernalillo was detected in 1995
- More consistent clusters are identified with trend-type detectors than step-type detectors

Cluster identified in 1994 with T-SCUSUM and  $\theta^* = 0.25$ 



Cluster identified in 1993 with T-SCUSUM and  $\theta^* = 0.5$ 



Cluster identified in 1994 with S-SCUSUM and  $\delta^* = 0.25$ 



Cluster identified in 1994 with S-SCUSUM and  $\delta^* = 0.5$ 



## Case Study 2: Covid 19 outbreaks in Miami, FL

- Weekly COVID-19 case counts observed from Feb to May 2020 in Miami, FL
  - Space-time CUSUM with trend shift was implemented for case counts observed weekly in zip codes
  - Two outbreaks are detected at two geographically distinct clusters



### Case Study 2:



## Conclusions

- Simulation study:
  - space-time CUSUM monitoring tuned for trend-type shifts can significantly outperform the counterparts tuned for sustained shifts in terms of the changepoint estimation accuracy (MSE)
  - a practical impact: accurate change-point estimates would enable health professionals to more accurately identify and isolate the disease emergence location and time and to devise more effective epidemic containment and mitigation policies
- Case studies:
  - Thyroid cancer data: trend-type space-time CUSUM gives more consistent cluster estimates regardless of tuning
  - Miami COVID data: high resolution monitoring for zip code cases allows taking more community focused containment measures
  - Identified clusters can be useful to identify gatherings with inadequate social distancing and implement targeted community testing. E.g., New York City's Health Department used surveillance to detect COVID-19 percent test positivity clusters and formulate containment solutions (Greene et al., 2020)



FAMU-FSU College of Engineering



Nour Alawad Email: <u>na18x@my.fsu.edu</u>

Arda Vanli

Email: <u>oavanli@eng.famu.fsu.edu</u>

