

Bayesian Hierarchical Modeling for Large-Scale Sensors Deployments and Applications

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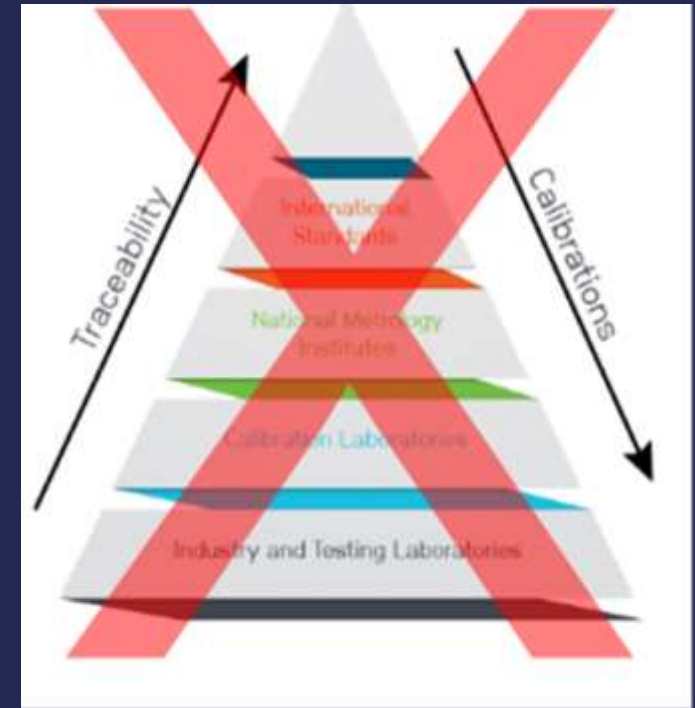
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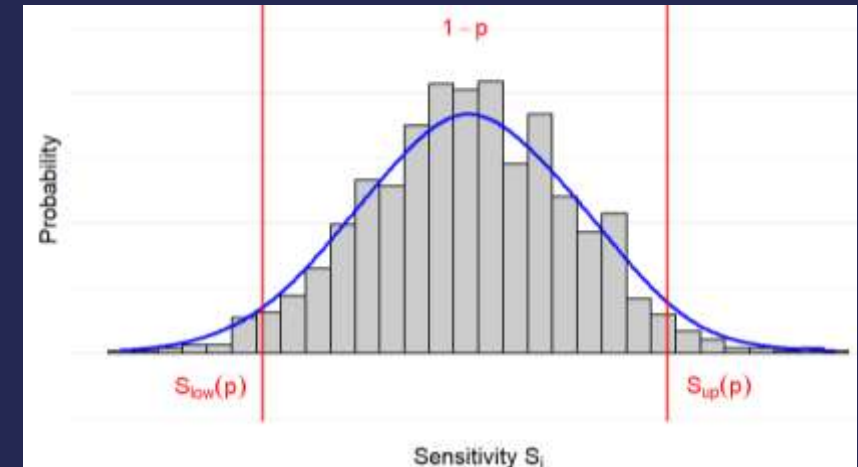
Introduction

- Millions of low-cost sensors being produced weekly:
 - impossibility to apply standard calibration procedures
 - large-scale calibration methods need to be introduced
- **Statistical methods:**
 - in-lab calibration of a **small sample** and **virtual calibration** of the entire batch
 - number of required experimental (“in-lab”) calibrations reduced
 - **CONS:** increased uncertainty and less reliability



Virtual Calibration

- Exploiting data from a reference, fully in-lab calibrated lot of N sensors
 - mixture distribution of N $\text{Normal}(S_i, u(S_i))$ modeling the sensitivities of the reference lot
 - determination of an interval of acceptable sensitivity values $[S_{\text{low}}(p), S_{\text{up}}(p)]$
- For any new lot from the same production, in-lab calibration of a small sample of sensors $n \ll N$
- Bayesian inference on the remaining $N-n$ sensors
- Evaluation of the reliability and estimation of sensitivity



A Bayesian model

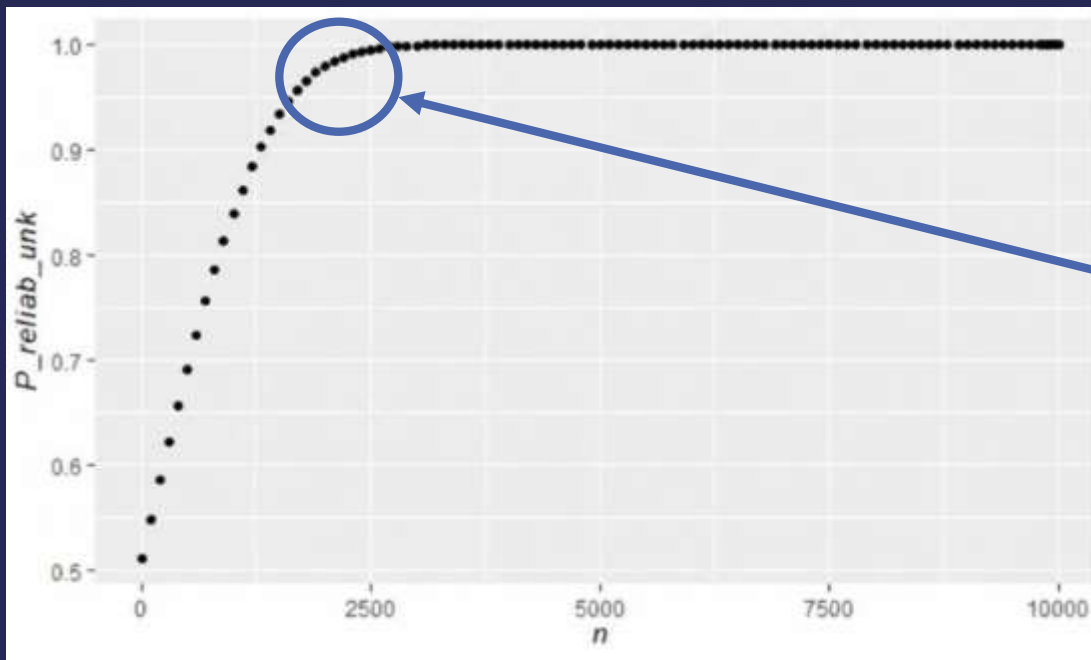
- Binomial prior to model the number of defective sensors C in the whole lot
- Hypergeometric likelihood on the number of defective items k in the sample of size n
- Conjugate model: Binomial posterior with updated parameters

$$f_{\text{post}}(C; k, n, N, p) = p^{C-k} (1-p)^{N-n-C+k} \binom{N-n}{C-k}$$

A metric for the reliability

A metric for batch reliability is defined based on posterior cumulative probability:

$$P_{\text{reliab,unk}} = \mathbb{P}_{\text{post}} = \sum_{C=k}^{C_{\text{bench}}} p^{C-k} (1-p)^{N-n-C+k} \binom{N-n}{C-k}$$



95 % of reliability with:

- at least $n = 1630$
- max $k = 49$

Figure. Variation of the probability $P_{\text{reliab,unk}}$ as a function of n

Hierarchical extension

- Introduction of a Beta hyperprior on the probability p of defective sensors
 - direct integration of prior industrial knowledge
 - explicit control over parameter variability
 - improvement in flexibility and robustness of the reliability
- CRUCIAL STEP: elicitation of hyperprior parameters to ensure the model coherently and accurately reflects prior knowledge

$$f_{\text{post}} = \binom{N-n}{C-k} \frac{\Gamma(C + \alpha) \Gamma(N - C + \beta)}{\Gamma(\alpha + \beta + N)} \frac{\Gamma(\alpha + \beta + n)}{\Gamma(\alpha + k) \Gamma(n - k + \beta)}$$

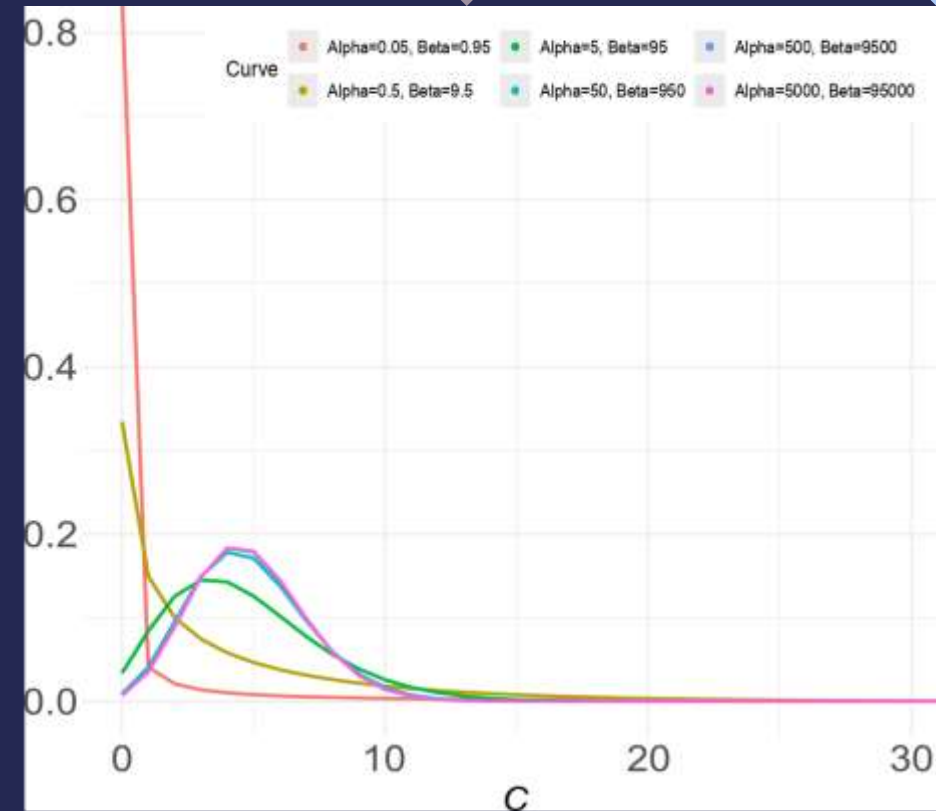
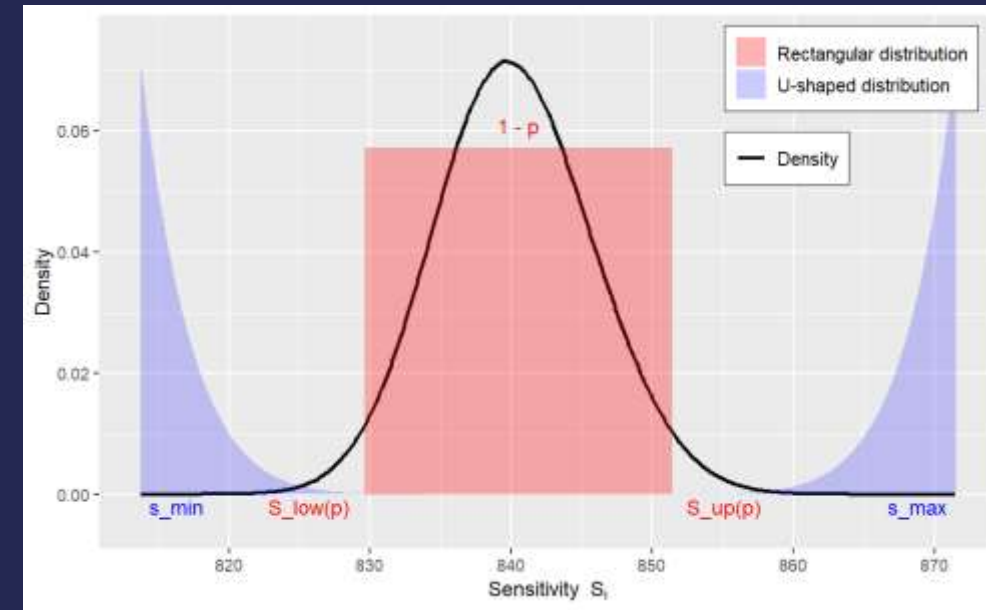


Figure. Comparison of the posterior distribution for different Beta hyperprior

Sensitivity estimate and uncertainty evaluation

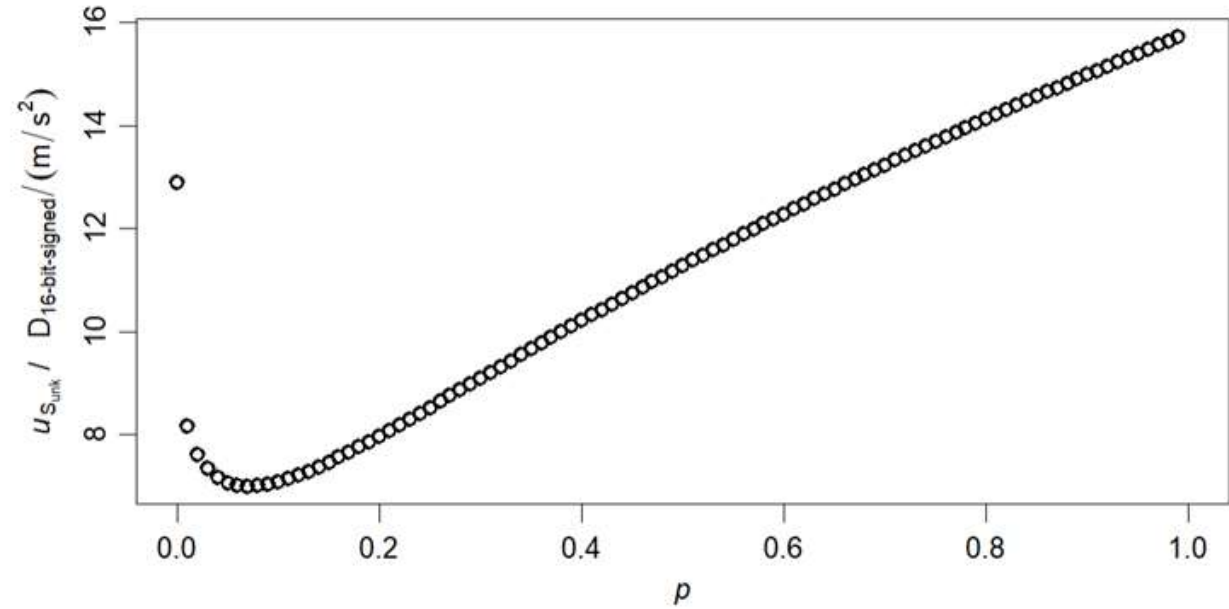
- Mixture distribution to model the sensitivity:
 - $[S_{low}(p), S_{up}(p)]$: limits of an interval encompassing the $(1 - p)$ fraction of acceptable sensors
 - $[S_{min}, S_{max}]$: maximum and minimum of the distribution
- S_{unk} is the variable to be estimated, arises from a mixture of:
 - a rectangular distribution on $[S_{low}(p), S_{up}(p)]$
 - a U-shaped distribution on $[S_{min}, S_{max}]$
- Weights of the mixture:-
 - $1 - \frac{C-k}{N-n}$: fraction of 'good' sensors
 - $\frac{C-k}{N-n}$: fraction of defective sensors



- Sensitivity:

$$S_{\text{unk}}(p) = \frac{S_{\text{low}}(p) + S_{\text{up}}(p)}{2}$$

$u(S_{\text{unk}})$ as a function of p



- Uncertainty:

$$u(S_{\text{unk}}) = \sqrt{\sum_{D=0}^{N-n} p_d \left[\left(1 - \frac{D}{N-n}\right) \frac{(S_{\text{up}}(p) - S_{\text{low}}(p))^2}{12} + \frac{D}{N-n} \frac{(S_{\text{max}} - S_{\text{min}})^2}{8} \right]}$$

where $p_d = \mathbb{P}(D = d)$

Figure. Value of the uncertainty as a function of p

Case study

- 100 nominally identical MEMS calibrated at INRiM
- Dataset randomly split: 50 benchmark – 50 unknown batch

- Acceptance interval:

$$[S_{\text{low}}, S_{\text{up}}] = [833.18, 849.14] \text{ (for } p = 0.06\text{)}$$

- Wider range for uncertainty evaluation:

$$[S_{\text{min}}, S_{\text{max}}] = [827.69, 855.06]$$

- Estimate of the sensitivity: $S_{\text{unk}} = 841.16 \text{ D}_{16\text{-bit-signed}} / (\text{ms}^{-2})$
- $n = 5, k = 0$ sensors counted outside acceptance interval
- True number of defective sensors: $C = 0$

Prior model	Mean	Mode	Reliability	$u(S_{\text{unk}})$
Beta(0.06, 0.94)	0.4	0	0.959	4.684
Beta(0.6, 9.4)	1.8	0	0.821	4.910
Beta(6, 94)	2.6	2	0.730	5.035
Beta(60, 940)	2.7	2	0.717	5.054
Original Model	2.7	2	0.717	5.056

Conclusion

- Industrial impact: drastically reduced times and costs for sensors network deployment while maintaining traceability
- Hierarchical modeling insight: better captures prior information structure, but highly sensitive to hyperparameter choice; weakly informative priors may lead to high reliability while increasing posterior variance
- A decision framework employing utility functions is under development to provide optimal sampling plans while keeping into account posterior variance

References

[1] A. Prato, F. Pennechi, G. Genta and A. Schiavi, A Bayesian statistical method for large-scale MEMS-based sensors calibration, Metrologia 61 (2024) 015005

[2] Prato A, Mazzoleni F and Schiavi A, Traceability of digital 3-axis MEMS accelerometer: simultaneous determination of main and transverse sensitivities in the frequency domain, Metrologia 57 (2020) 035013

[3] Ballario A., Prato A., Schiavi A., Pennechi F., Bayesian Hierarchical Modelling for Large-Scale Sensor Calibration, submitted to Quality and Reliability Engineering International, under review

**Thank you for
the attention!**
